Contents:

Dimensionality of interpersonal curiosity

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Abstract

Interpersonal curiosity (IPC) is the desire for new information about people. Fifty-one IPC items were administered to 321 participants (248 women, 73 men), along with other measures of curiosity and personality. Three factors were identified from which five-item subscales were developed that had good internal consistency: *Curiosity about Emotions*, *Spying and Prying*, and *Snooping*. Confirmatory factor analysis indicated the three-factor model had acceptable fit. The IPC scales correlated positively with other curiosity measures and interest in gossip, providing evidence of convergent validity. Divergent validity was demonstrated in finding the other curiosity scales correlated more highly with each other than with IPC; parallel results were found for the gossip measures.
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1. Introduction

Curiosity is the intrinsic desire for new information that will stimulate interest or relieve uncertainty (Litman, 2005). Historically, researchers have investigated curiosity for three types of information: intellectual knowledge (Berlyne, 1954; Litman & Spielberger, 2003), sensory stimulation

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(Berlyne, 1957, 1958; Collins, Litman, & Spielberger, 2004), and experiences characterized as adventurous or thrilling (Litman, Collins, & Spielberger, 2005; Zuckerman, 1994). However, there is another kind of information that has received considerably less attention by curiosity researchers – information about *people*.

People-information differs from other kinds of knowledge in several ways: First, people-information is extraordinarily complex (Fiske, 1995), and includes knowledge about individuals' experiences, their public and private behaviors, and also their thoughts and feelings. Second, people-information often has special value in the social world; obtaining it is important for social comparison (Festinger, 1954; Snyder & Ickes, 1985), while its transmission (i.e., gossiping) plays a role in forming friendships (Rosnow, 2001) and attacking adversaries (Galen & Underwood, 1997). Additionally, because people-information often involves private matters that cannot be easily confirmed, individuals differ in their willingness to share it (Nevo, Nevo, & Derech-Zehavi, 1994), as well as their attitudes about its reliability as a source of information (Litman & Pezzo, 2005).

The intrinsic motive to seek people-information is referred to as *interpersonal curiosity* (IPC), for which the signature measure is the interpersonal curiosity scale (IPCS; Singer & Antrobus, 1963). The IPCS is a 12-item instrument that assesses differential tendencies to wonder about people's public and private lives. In past research, the IPCS has demonstrated satisfactory internal consistency ($\alpha > 0.80$), and evidenced construct validity on the basis of positive correlations with gossip tendencies and negative correlations with social approval needs (Litman & Pezzo, 2005).

While the IPCS appears to be a reliable and valid measure of individual differences in IPC, we note two important limitations of this instrument: First, the IPCS items inquire almost exclusively about interest in people's *external* life-experiences (e.g., what they do for a living); there are no items that assess curiosity about *internal* life-experiences such as thoughts or feelings, which are significant sources of people-information (Fiske, 1995); being motivated to seek information about people's internal experiences may play an important role in the development of abilities such as empathy (Håkansson & Montgomery, 2003) and emotional intelligence (Mayer & Salovey, 1990). Second, the IPCS items all refer to passively wondering about other people, but do not address active information-seeking behaviors such as asking questions or spying,[1] which may be influenced by social approval constraints or experiences of anxiety (Trudewind, 2000).

Given that the IPCS does not assess differential interests in learning about internal life-experiences or engaging in active information-seeking, these two important aspects of IPC remain unmeasured. Thus, the goal of the present study was to develop a new instrument for assessing individual differences in facets of IPC not previously taken into account. Moreover, given that the structure of curiosity instruments reflects both the type of information sought as well as the methods employed in seeking it (e.g., Litman & Spielberger, 2003), the dimensionality of IPC was also examined, with the expectation that emergent factors would correspond to different kinds of people-information (i.e., internal vs. external experience) and different methods of attaining it (i.e., overt vs. covert), from which subscales could be constructed. The internal consistency of the

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[1] Renner (2006) recently developed a similar scale that differentiates between covert and overt social exploration, which was not available for the present study.

newly developed IPC scales as well as the IPCS was evaluated to draw comparisons between these instruments.

We hypothesized that the IPC scales would demonstrate convergent validity by correlating positively with measures of curiosity for other types of information, and also to scales that involved sharing and valuing people-information in the form of gossip. Divergent validity for the IPC scales would be demonstrated by finding stronger correlations between IPC and gossip scales as compared to the correlation between the other curiosity scales and gossip measures. Moreover, we expected measures of curiosity for information unrelated to people would correlate more highly with each other than with IPC; we expected similar findings for the gossip measures.

We also expected the new IPC measures to show divergence from sociability, given that IPC does not necessarily involve a desire to interact with others more than required to obtain new information about them – like other forms of curiosity, we do not consider IPC derivative of other motives (Montgomery & Monkman, 1955), such as extraversion. Rather, those who are interpersonally curious are assumed to be intrinsically motivated to obtain people-information primarily for the sake of learning something interesting or eliminating uncertainty.

Finally, in keeping with recent research findings by Renner (2006), we expected aspects of IPC that involved covert information-seeking would correlate positively with anxiety but negatively with social approval needs, whereas, overt methods were expected to correlate negatively with anxiety and positively with social approval.

2. Method

2.1. Participants

Undergraduate students were recruited from psychology courses at a large urban university. Data were collected from two samples over consecutive semesters:

Sample 1 (fall): 324 participants (233 women, 91 men), ranging in age from 18 to 58 ($M = 21.94$; SD = 6.42).
Sample 2 (spring): 229 participants (159 women, 70 men), ranging in age from 18 to 51 ($M = 22.71$; SD = 6.52).

2.2. Measures and procedures

IPC item pool. We constructed 51 face-valid items that expressed a desire for new information about people, including their life-experiences, thoughts, and feelings. A number of items also referred to seeking private information. To ensure a broad range of possible items for measuring interest in people-information, we also included items that indicated self-determination in obtaining information about others as well as confidence that a positive result would follow the expenditure of effort to obtain such information (see Table 1 for examples). Given that expressions of interpersonal curiosity are theorized to be highly positive (Kashdan & Roberts, 2004), and people-information is considered interesting across cultures (Brenneis, 1984; Cox, 1970), the wording of

Table 1
Principal axis factor loadings with oblique rotation of the 43 IPC items ($N = 334$)

	I	II	III
Attend to non-verbal messages	**0.72**	−0.11	−0.02
Observe people's expressions to figure out how they feel	**0.65**	−0.18	0.14
Imagine what people are thinking from expressions	**0.65**	−0.02	0.19
Try to understand people's feelings	**0.65**	−0.22	0.13
People open up to me about how they feel	**0.61**	0.08	−0.11
Others tell me things they wouldn't tell	**0.60**	0.16	−0.07
Make an effort to interpret feelings from voice	**0.56**	−0.08	0.18
People confide secrets to me	**0.56**	0.29	−0.11
Try to figure out what others are feeling by looking	**0.53**	−0.04	0.14
Try and figure out people's motives	**0.50**	0.23	0.04
Good at getting people to open up	**0.43**	**0.33**	−0.04
Someone is hiding something, I find out	**0.40**	**0.32**	−0.12
Another person's experience; wonder how it feels	**0.37**	0.02	0.21
Private lives are fascinating	−0.02	**0.67**	0.14
Someone's diary would read it	−0.11	**0.66**	−0.02
Wish I could turn invisible to spy	−0.07	**0.65**	−0.03
Interested in private lives	−0.05	**0.53**	0.20
Find myself eavesdropping	−0.04	**0.52**	0.10
Think about being an investigative reporter	0.14	**0.52**	−0.03
Like to know secrets	−0.08	**0.51**	0.16
Feel comfortable asking about people's private lives	0.20	**0.50**	−0.11
Find myself looking over people's shoulder	−0.21	**0.47**	0.25
Think about interviewing others	0.11	**0.46**	−0.05
Can persuade people to tell secrets	**0.37**	**0.45**	−0.10
Would make a good detective	0.25	**0.42**	−0.02
Curious to know what people are talking about	0.16	**0.40**	0.23
Try to figure out what makes people tick	0.18	**0.39**	0.12
Would enjoy reading confessions	−0.09	**0.35**	**0.31**
Wonder what people think of me	0.04	**0.30**	0.21
Like to look at things in people's rooms	0.01	0.05	**0.61**
Love going into people's houses to see how they live	0.04	0.02	**0.60**
Like to see how others live	0.01	0.07	**0.58**
Wonder what other people's interests are	0.05	0.03	**0.57**
Like to know what other people do	0.05	0.09	**0.54**
Wonder what people do or don't know	0.00	0.06	**0.51**
Like to know music people listen to	0.17	−0.06	**0.50**
Shuffle through things because intrigued	0.02	0.15	**0.44**
Wonder what people's lives were like	0.15	0.01	**0.40**
Like to know what people are reading	−0.04	0.13	**0.39**
Wonder who people are friends with	0.09	−0.03	**0.38**
Curious how people view the world	0.21	−0.02	**0.32**
Wonder what people look at on computer	−0.16	**0.30**	**0.31**
Like to look at pictures of my friends	0.19	−0.01	0.29

Note: Items in italics were selected for the IPC scale.

the items primarily referred to the pleasurable stimulation of interest rather than the reduction of unpleasant feelings of uncertainty (Litman, 2005). For each item, participants indicated how they

"generally feel" by rating themselves on a four-point frequency scale ranging from "Almost Never" to "Almost Always".

Interpersonal Curiosity Scale (IPCS; Singer & Antrobus, 1963) consists of 12-items ($\alpha > 0.80$) that assess interest in learning about others. Participants indicated how much each item was true for them using a five-point scale anchored by "Definitely not true" and "Very true."

Epistemic Curiosity scale (EC; Litman & Spielberger, 2003) is a 10-item instrument (α range = 0.79–0.86) that measures positive feelings of interest in gaining intellectual knowledge not directly related to people, such as learning new ideas and how things work. The EC scale uses the same rating format as the IPC item pool.

Curiosity as a feeling-of-deprivation scale (CFD; Litman & Jimerson, 2004) is a 15-item instrument (α range = 0.84–0.88) that, like the EC scale, also assesses the desire to obtain intellectual knowledge unrelated to people (e.g., obtain facts, solve problems). However, unlike the EC scale, the CFD scale emphasizes acquiring knowledge to reduce feelings of tension due to uncertainty. The CFD scale also uses the same rating format as the IPC item pool.

Attitudes Towards Gossip scale (ATG; Litman & Pezzo, 2005) is a 12-item measure ($\alpha \geqslant 0.80$) that assess attitudes about the social and moral value of gossip. Participants responded to each item using a five-point scale, anchored by "Disagree Strongly" and "Agree Strongly."

Tendency to Gossip Questionnaire (TGQ; Nevo et al., 1994) is comprised of 20 items that assess tendencies to talk about others in relation to a variety of different topics. Each TGQ item is rated on a seven-point frequency scale ranging from "Never" to "Always."

International Personality Item Pool Extraversion Scale (IPIP-Ext; Goldberg, 1999) is a 10-item measure ($\alpha = 0.86$) of this Big Five construct. Participants indicated how well each item described themselves using a five-point scale ranging from "Very Inaccurate" to "Very Accurate".

Social Desirability Scale (SDS; Strahan & Gerbasi, 1963) is a 20-item measure (K–R range = 0.73–0.83) of social approval needs. Respondents indicate whether each item statement is either "true" or "false".

Trait Anxiety Scale of the State Trait Personality Inventory (T-Anx; Spielberger, 1979) is a 10-item measure ($\alpha = 0.80$) of individual differences in anxiety, which asks respondents to report how frequently corresponding emotional states are generally experienced. Participants rated themselves on the same scale used for the IPC items.

2.3. Procedure

The experimenter introduced himself to the participants, indicated that the goals of the study were to learn about their feelings and attitudes, and that additional information would be provided after participating. Approximately 30 min were required to participate.

3. Results

In order to develop an internally consistent IPC scale, items with corrected item-total (CIT) correlations <0.30 and mean inter-item correlations <0.20 were dropped (Comrey, 1988), resulting in the deletion of eight items. To evaluate the dimensionality of IPC, responses to the remaining 43 items by sample 1 were analyzed in an iterated principal-axis factor analysis with oblique

(promax) rotation, using the squared multiple correlation as the communality estimate.[2] Examination of the scree plot and parallel analyses of the eigenvalues (9.78, 3.18, 1.97, 1.35, 1.25, 0.93, ...) suggested that three or four factors should be extracted.

Rotated loadings for the three-factor solution are reported in Table 1. Of the 43 IPC items, 39 had dominant salient loadings ($\geqslant$0.30) on a single factor, of which five had salient dual loadings. Factor I was defined by 11 items with dominant loadings $\geqslant$0.37 and no salient dual loadings. These items involved *Curiosity about Emotions* (CE), and reflected a desire to learn people's feelings. On Factor II, 14 items with loadings $\geqslant$0.36 and no dual loadings expressed willingness to engage in *Spying and Prying* (SP). Twelve items with dominant loadings $\geqslant$0.32 on Factor III described being curious about people's interests and life-experiences, and involved *Snooping* (Sn) for information by prowling about or investigating their belongings. The factors were moderately correlated, with Factors II and III overlapping the most ($r_{12} = 0.29$; $r_{13} = 0.35$; $r_{23} = 0.46$).

When four factors were extracted, the factor pattern was highly similar to that of the three-factor solution, with one notable difference: The third factor split into two weaker factors, each of which was difficult to interpret as meaningfully distinct from the other. Given the similar composition of the factors, but reduced interpretability, the four-factor solution was not considered further.

The next step was to select items for measuring each factor on strength of loadings and content. From Factor I, of the 11 dominant items, six were eliminated because of content that was considered redundant with items that had better loadings. The same procedure was applied in selecting seven items from the second factor. An eighth item that referred to "eavesdropping" was dropped because its content suggested passively absorbing information rather than actively seeking it. Five items were similarly selected from Factor III. These 17 items (italicized in Table 1) formed preliminary subscales for measuring each factor.

Responses to the 17 IPC items by sample 2 were submitted to confirmatory factor analysis using maximum likelihood estimation. Several goodness of fit indices were examined, including chi-square (χ^2), comparative fit index (CFI), non-normed fit index (NNFI), McDonald's fit index (MFI), root mean square error of approximation (RMSEA), and expected cross-validation index (ECVI). Because χ^2 values can become inflated with large samples (Raykov, 1998; Russell, 2002), models with χ^2 to *df* ratio <2.0 were considered acceptable (Hatcher, 1994). Values $\geqslant$0.95 for CFI and NNFI and $\leqslant$0.90 for MFI indicate close fit (Hu & Bentler, 1999), while values $\geqslant$0.90 are acceptable (Raykov, 1998). For RMSEA, values $\geqslant$0.06 indicate close fit (Hu & Bentler, 1999). Although there is no specific cutoff for ECVI, lower values indicate superior fit (Hatcher, 1994).

As shown in Table 2, χ^2 was significant for the 17-item model ($p < .001$), while the χ^2:*df* = 2.0. CFI, NNI, and MFI were <0.90, and RMSEA was >0.06 indicating poor model fit. Two items ("someone's diary, would read it"; "private lives are fascinating") that were assigned to factor II were complex indicators with large Lagrange multipliers; these two items were retained for the total scale, but deleted from the model and subscales. As may be noted in Table 2, the revised model demonstrated improved fit. Although χ^2 was still significant, χ^2:*df* was now <2.0; CFI was

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[2] Although the composition of men in each sample was not considered large enough to conduct meaningful analyses of gender differences, several exploratory factor analyses (not reported) indicated the same factors for both sexes. Given these preliminary findings, the final analyses were conducted on the combined sample. More analysis of possible gender differences in IPC will be important for future research.

Table 2
Goodness of fit (GOF) indices for the three-factor IPC model ($N = 229$)

GOF index	17-Items	15-Items
$\chi^2(df)$ $[\chi^2{:}df]$	232.75 (116) [2.00]	128.50 (87) [1.48]
CFI	0.886	0.953
NNFI	0.887	0.944
MFI	0.766	0.910
RMSEA [95% CI]	0.071 [0.06–0.08]	0.047 [0.03–0.06]
ECVI [95% CI]	1.44 [1.25–1.66]	0.92 [0.79–1.08]

Both χ^2 values are significant, $p < 0.01$.

>0.95, MFI and NNI were >0.90, and RMSEA was <0.06. In comparing fit between the first and second model, the ECVI was lower for the 15-item model. The final three factor model is presented in Fig. 1; all loadings ($Mdn = 0.64$) were significant ($p < 0.001$).

Means and standard deviations for the 17-item IPC total scale and its three five-item subscales are reported in Table 3. Alpha coefficients, SE and 95% CI for alpha, mean inter-item correlations, and mean CIT correlations for these instruments are also reported in this table. These data are also reported for the IPCS for comparison. Alphas for the IPC scale were fairly high ($\alpha = 0.82$). Although alphas were lower for the subscales, the mean inter-item correlations were very high, ranging from 0.46 to 0.57, suggesting that the lower alphas were due in part to these
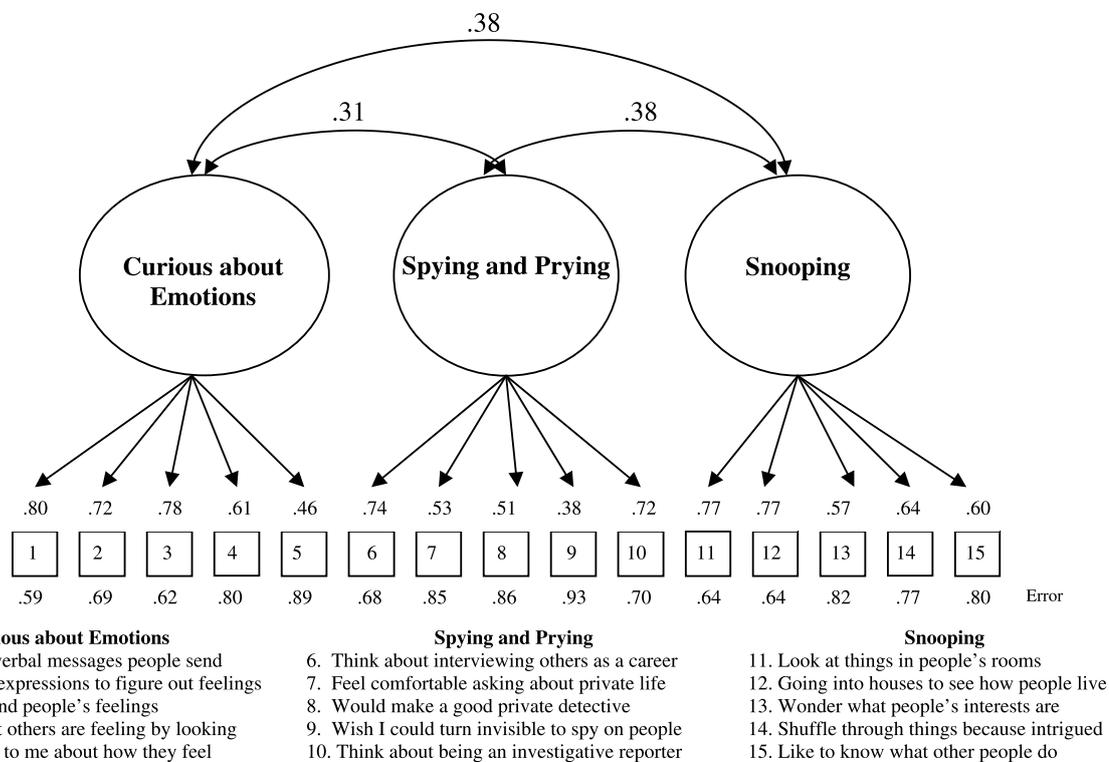

Fig. 1. Path diagram of the three-factor structural IPC model ($n = 229$).

Table 3

Means, standard deviations, reliability indices, and correlations among IPC measures for sample 1 ($N = 324$) and sample 2 ($N = 229$)

		M (SD)	α (SE$_\alpha$)	95% CI$_\alpha$	M inter-item r	M CIT	1	2	3	4
1.	IPC Total	41.61 (7.65)	0.82 (0.015)	0.79–0.85	0.26	0.41				
		41.10 (7.98)	0.82 (0.018)	0.78–0.85	0.26	0.42				
2.	IPC-CE	14.59 (3.03)	0.76 (0.021)	0.72–0.80	0.52	0.53	0.63			
		14.83 (3.29)	0.81 (0.020)	0.77–0.85	0.57	0.60	0.65			
3.	IPC-SP	10.45 (3.40)	0.73 (0.024)	0.68–0.78	0.48	0.50	0.77	0.23		
		10.17 (3.41)	0.71 (0.030)	0.65–0.77	0.46	0.47	0.74	0.26		
4.	IPC-Sn	12.32 (3.05)	0.74 (0.022)	0.70–0.79	0.49	0.51	0.73	0.28	0.32	
		11.96 (3.31)	0.79 (0.021)	0.75–0.84	0.55	0.57	0.75	0.31	0.31	
5.	IPCS	36.91 (5.03)	0.64 (0.035)	0.58–0.71	0.13	0.29	0.23	0.08	0.16	0.20
		36.94 (6.04)	0.65 (0.034)	0.56–0.72	0.21	0.30	0.43	0.21	0.28	0.39

r's $\geqslant 0.16$ are significant, $p < 0.05$.

For each scale, first row = sample 1, second row = sample 2.

measures having fewer items rather than a lack of internal consistency (Clark & Watson, 1995; Cortina, 1993). SE for the alphas were low (<0.1), suggesting that the IPC scale items were relatively homogeneous. SE tended to be somewhat higher for the subscales relative to the total scale, again due primarily to having fewer items (Duhachek & Iacobucci, 2004). Mean CIT correlations for the IPC scales were all moderately strong ($\geqslant 0.40$). As would be expected, the IPC scale correlated very highly with its subscales. Alphas for the IPCS were quite low in both samples, as were mean inter-item and CIT correlations. The IPC scales had small positive correlations with the IPCS in both samples ($Mr_{s1} = 0.18$; $Mr_{s2} = 0.30$).

Pearson correlations between the IPC scales, the IPCS, and the other personality traits for both samples are reported in Table 3. As hypothesized, the IPC scales correlated positively with the two measures of intellectual curiosity, demonstrating convergent validity. Finding the two intellectual curiosity scales were more highly correlated with each other than with the IPC scales indicated divergence of IPC from curiosity for other forms of information. Although the IPC scales appeared to be about equally related to both the EC and CFD scales, regression analyses revealed this relationship was due primarily to correlations with the CFD scale (*sample 1:* $\beta_{CFD} = 0.29$, SE = 0.08, $t = 3.42$, $p < .001$; $\beta_{EC} = 0.23$, SE = 0.12, $t = 1.92$ n.s; *sample 2:* $\beta_{CFD} = 0.38$, SE = 0.11, $t = 3.48$, $p < .001$; $\beta_{EC} = 0.03$, SE = 0.15, $t = 0.17$ n.s). That the IPC scale was generally more related to CFD was contrary to expectations. The IPCS was only weakly and inconsistently related to either intellectual curiosity scale, which raises questions about what the IPCS measures.

As hypothesized, positive correlations were found between the IPC scales and the TGQ, providing additional evidence of convergent validity. Unexpectedly, the IPC scales were essentially unrelated to the ATG, possibly because the ATG measures attitudes about the value of people-information rather than interest in sharing it. Consistent with previous research, the IPCS was positively correlated with both gossip scales. As expected, the IPC scales were more highly related to gossip, as measured by the TGQ scale, than the two other curiosity measures were. However, the small negative correlations between the two other curiosity scales and the ATG was unanticipated. As expected, the two gossip instruments were more highly correlated with each other than

Table 4
Means, standard deviations, reliability indices, and correlations among IPC measures and other personality traits for sample 1 ($N = 324$) and sample 2 ($N = 229$)

	M (SD)	α	IPC Total	IPC-CE	IPC-SP	IPC-Sn	IPCS	EC	CFD	Extra	TGQ	ATG	SDS
EC	26.16 (5.96)	0.85	0.44	0.35	0.35	0.27	0.17						
	25.74 (6.26)	0.86	0.38	0.38	0.27	0.21	0.11						
CFD	39.31 (8.50)	0.88	0.48	0.37	0.41	0.25	0.20	0.77					
	39.17 (8.75)	0.89	0.43	0.41	0.28	0.27	0.06	0.85					
Extraversion	34.21 (6.79)	0.82	0.16	0.26	−0.03	0.09	0.07	0.18	0.14				
	33.70 (6.89)	0.80	0.11	0.18	0.08	−0.10	0.04	0.19	0.17				
TGQ	53.38 (12.92)	0.89	0.36	0.14	0.30	0.26	0.42	0.20	0.24	0.25			
	52.82 (13.03)	0.89	0.35	0.16	0.27	0.28	0.44	0.01	0.05	0.04			
ATG	33.37 (7.04)	0.73	0.08	−0.18	0.13	0.08	0.23	−0.11	−0.10	−0.05	0.40		
	31.90 (7.37)	0.76	0.12	−0.11	0.16	0.14	0.25	−0.23	−0.20	−0.01	0.45		
SDS	9.74 (3.33)	0.66	0.12	0.16	0.00	0.12	−0.15	0.05	0.04	−0.01	−0.28	−0.22	
	10.03 (3.32)	0.66	−0.09	0.01	−0.04	−0.10	−0.20	0.15	0.10	−0.02	−0.31	−0.26	
T-Anx	20.63 (6.20)	0.86	0.16	−0.10	0.25	0.08	0.09	0.06	0.20	−0.35	0.11	0.05	−0.08
	20.89 (6.01)	0.83	0.23	0.01	0.17	0.27	0.12	−0.08	0.02	−0.44	0.16	0.10	−0.19

r's ⩾ 0.16 are significant *p* < .05.
For each scale, first row = sample 1, second row = sample 2.

with the IPC instruments, demonstrating divergence of IPC from gossip. The IPC scales and the IPCS showed divergence from extraversion, by correlating essentially zero, with the exception of the CE subscale, which had small positive correlations.

Contrary to expectations, the IPC scales were essentially unrelated to social desirability, whereas, the IPCS had very small negative correlations with the SDS. As expected, the SP and Sn subscales, which involve somewhat more covert forms of information-seeking, had small, but significant positive correlations with T-anx (Study 1); the IPCS was uncorrelated with anxiety (see Table 4).

4. Discussion

Three IPC factors were identified: curiosity about emotions (CE), spying and prying (SP), and snooping (Sn). As hypothesized, these factors reflected different sources of people-information. Seventeen items were selected to form an internally consistent IPC scale with subscales to assess each factor. With the omission of two items retained for the total scale, but not included in the subscales, acceptable fit was achieved.

As hypothesized, the IPC scale was correlated with both intellectual curiosity scales, providing evidence of convergent validity. Contrary to expectations, the IPC scale was more related to deprivation-type than interest-type curiosity, suggesting that most of the pleasure derived from learning about others is due to uncertainty reduction (Litman, 2005). Also as hypothesized, the IPC scales were positively related to sharing gossip, providing further evidence of convergent validity.

Divergence of IPC from other measures of curiosity was demonstrated by finding stronger correlations between the IPC scales and TGQ as compared to the relationship between this gossip scale and the two other curiosity measures. Divergent validity was also evidenced in finding the two other curiosity scales correlated more highly with each other than with IPC; parallel results were found for gossip. These findings suggest that those who are interpersonally curious want new information about people in order to reduce uncertainty about them. Moreover, this motive may be differentiated from wanting to learn intellectual knowledge and also from interest in sharing people-information with others.

Generally consistent with our hypothesis, except for the CE subscale, the IPC scales were uncorrelated with extraversion. This finding is in line with the theoretical view that IPC is primarily a motive to acquire information, and is not derivative of other motives or needs (Montgomery & Monkman, 1955). The small positive correlations with the CE subscale may indicate that being curious about people's feelings is more likely to motivate social interaction, perhaps because obtaining information about the emotions of others requires direct interaction (e.g., to examine facial expressions).

Positive correlations were found for the SP and Sn subscales with anxiety, possibly because these methods of information-seeking are motivated in part by social anxiety (Renner, 2006). However, finding the IPC scales were not consistently correlated with social desirability suggests that wanting information about others has little to do with social approval needs. Another possibility for these relationships is that individuals may be more likely to be concerned with seeking people-information surreptitiously when worried about it how it may affect them.

In future research it will be important to examine relationships between individual differences in each IPC dimension and relevant information-seeking behaviors such as tending to inquire about people's feelings (CE) or their daily activities (Sn), and willingness to pry into the affairs of others (SP). It will be especially important to investigate how differential tendencies to engage in such information-seeking behaviors interact with social skills and people's ability to understand the needs and intentions of others.

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